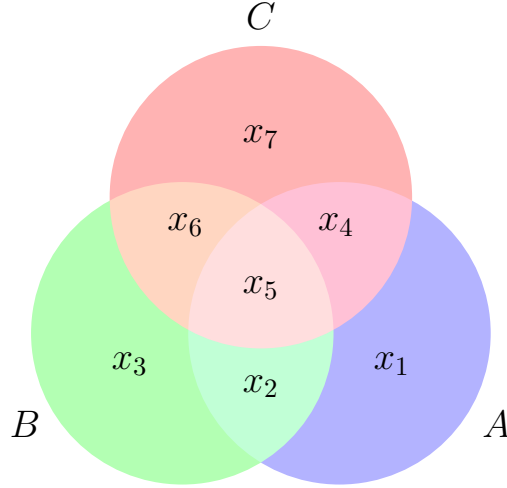


## ANSWER SHEET 1

**Assignment 1.** Define  $x_1 = P(A \cap (B \cup C)^c)$ ,  $x_2 = P(A \cap B \cap C^c)$ ,  $x_3 = P(B \cap (A \cup C)^c)$ ,  $x_4 = P(A \cap C \cap B^c)$ ,  $x_5 = P(A \cap B \cap C)$ ,  $x_6 = P(B \cap C \cap A^c)$ ,  $x_7 = P(C \cap (A \cup B)^c)$ .



Then, using the information given, we have

$$P(A) = x_1 + x_2 + x_4 + x_5 = 0.4 \quad (1)$$

$$P(B) = x_2 + x_3 + x_5 + x_6 = 0.7 \quad (2)$$

$$P(A \cup B) = x_1 + x_2 + x_3 + x_4 + x_5 + x_6 = 0.8 \quad (3)$$

$$P(C \cap (A \cup B)) = x_4 + x_5 + x_6 = 0.2 \quad (4)$$

$$P(B \cap (A \cup C)) = x_2 + x_5 + x_6 = 0.4 \quad (5)$$

$$P(C) = x_4 + x_5 + x_6 + x_7 = 0.3 \quad (6)$$

$$P(B \cap C) = x_5 + x_6 = 0.2 \quad (7)$$

Using (1)-(3), we get

$$P(A \cap B) = x_2 + x_5 = 0.3 \quad (8)$$

Using (8) and (5), we get  $x_6 = 0.1$ . Using (4) and (7), we get  $x_4 = 0$ . Putting these in (4), we get  $x_5 = 0.1$ . So, (6) yields  $x_7 = 0.1$ , and (5) yields  $x_2 = 0.2$ . From (1) and (2), we now get  $x_1 = 0.1$  and  $x_3 = 0.3$ .

- (a) The probability that exactly two of  $A$ ,  $B$  and  $C$  occur equals  $x_2 + x_4 + x_6 = 0.3$ .  
 (b) The probability that none of  $A$ ,  $B$  and  $C$  occur equals  $1 - \sum_{k=1}^7 x_k = 0.1$ .  
 (c) The probability that  $A$  and exactly one of  $B$  and  $C$  occur equals  $x_2 + x_4 = 0.2$ .

**Assignment 2.** Denote by ‘H’ and ‘T’ the events that a head and a tail occurs, respectively. A sample point  $\{x, y, z\}$  with  $x, y, z \in \{H, T\}$  will denote that  $x$  occurs for  $A$ ,  $y$  occurs for  $B$  and  $z$  occurs for  $C$ .

- (a)  $\Omega = \{\{H, H, H\}, \{H, H, T\}, \{H, T, H\}, \{H, T, T\}, \{T, H, H\}, \{T, H, T\}, \{T, T, H\}, \{T, T, T\}\}$ .  
 (b) Assuming the each coin is fair, all the sample points may be assigned equal probabilities, i.e.,  $P(\{x, y, z\}) = 1/8$  for all  $\{x, y, z\} \in \Omega$ .  
 (c)  $P(A \text{ wins on first toss}) = P(\{H, T, T\} \cup \{T, H, H\}) = 1/4$ . The probabilities are the same

for  $B$  as well as  $C$  winning on the first toss.

(d)  $P(\text{no winner on first toss}) = P(\{H, H, H\} \cup \{T, T, T\}) = 1/4$ .

(e) Let  $F$  and  $G$  denote the events that  $A$  wins on first toss and that the winner is decided on first toss, respectively. Then,  $P(G) = P(A \text{ wins on first toss}) + P(B \text{ wins on first toss}) + P(C \text{ wins on first toss}) = 3/4$ . Since  $F \subset G$ , we have  $P(F \cap G) = P(F) = 1/4$ . So,  $P(F | G) = P(F \cap G)/P(G) = 1/3$ .

*(Intuition : since all events are equally likely, if there is a winner on the first toss, then it has to be one of  $A$ ,  $B$  and  $C$  with equal probability. This probability is obviously  $1/3$ ).*

**Assignment 3.** Since only one prisoner will be pardoned and that too at random,  $P(A \text{ will be pardoned}) = P(B \text{ will be pardoned}) = P(C \text{ will be pardoned}) = 1/3$ .

However, as per the understanding between prisoner  $A$  and the jailer, the following are true :

(i)  $P(A \text{ is told that } C \text{ will be executed} | A \text{ will be pardoned}) = P(A \text{ is told that } B \text{ will be executed} | A \text{ will be pardoned}) = 1/2$  (assuming that the jailer chooses between  $B$  and  $C$  at random),

(ii)  $P(A \text{ is told that } C \text{ will be executed} | B \text{ will be pardoned}) = 1$  (since the jailer is left with no other choice and he cannot tell  $A$  his fate),

(iii)  $P(A \text{ is told that } C \text{ will be executed} | C \text{ will be pardoned}) = 0$  (assuming that the jailer does not lie).

Thus, by the law of total probability,

$$\begin{aligned} &P(A \text{ is told that } C \text{ will be executed}) \\ &= P(A \text{ will be pardoned}) \times P(A \text{ is told that } C \text{ will be executed} | A \text{ will be pardoned}) \\ &+ P(B \text{ will be pardoned}) \times P(A \text{ is told that } C \text{ will be executed} | B \text{ will be pardoned}) \\ &+ P(C \text{ will be pardoned}) \times P(A \text{ is told that } C \text{ will be executed} | C \text{ will be pardoned}) \\ &= (1/3) \times (1/2) + (1/3) \times 1 + (1/3) \times 0 = 1/2. \end{aligned}$$

(a) By Bayes' theorem,

$$\begin{aligned} &P(A \text{ will be pardoned} | A \text{ is told that } C \text{ will be executed}) \\ &= P(A \text{ will be pardoned} \cap A \text{ is told that } C \text{ will be executed})/P(A \text{ is told that } C \text{ will be executed}) \\ &= P(A \text{ will be pardoned}) \times P(A \text{ is told that } C \text{ will be executed} | A \text{ will be pardoned})/(1/2) \\ &= (1/3) \times (1/2)/(1/2) = 1/3. \end{aligned}$$

(b) By Bayes' theorem,

$$\begin{aligned} &P(B \text{ will be pardoned} | A \text{ is told that } C \text{ will be executed}) \\ &= P(B \text{ will be pardoned} \cap A \text{ is told that } C \text{ will be executed})/P(A \text{ is told that } C \text{ will be executed}) \\ &= P(B \text{ will be pardoned}) \times P(A \text{ is told that } C \text{ will be executed} | B \text{ will be pardoned})/(1/2) \\ &= (1/3) \times 1/(1/2) = 2/3. \end{aligned}$$

So, prisoner  $A$  was wrong in his assessment that the jailer's information will increase his chances of getting a pardon. In fact, the remaining prisoner is the one who benefits from this information (provided of course that  $A$  shares this information with him!).

(c) Using similar argument as above, we have  $P(A_j \text{ will be pardoned}) = 1/n$  for  $j = 1, 2, \dots, n$ , and

(i)  $P(A_1 \text{ is told that } A_3, A_4, \dots, A_n \text{ will be executed} | A_1 \text{ is pardoned}) = 1/(n-1)$  (assuming that the jailer chooses randomly among one of  $(n-1)$  other prisoners to exclude from his list that he will tell  $A_1$ ),

(ii)  $P(A_1 \text{ is told that } A_3, A_4, \dots, A_n \text{ will be executed} | A_2 \text{ is pardoned}) = 1$  (since the jailer has no other option in this case),

(iii)  $P(A_1 \text{ is told that } A_3, A_4, \dots, A_n \text{ will be executed} \mid A_k \text{ is pardoned}) = 0$  for all  $k = 3, 4, \dots, n$  (assuming that the jailer does not lie).

Thus, by the law of total probability,

$$P(A \text{ is told that } A_3, A_4, \dots, A_n \text{ will be executed}) \\ = (1/n) \times [1/(n-1)] + (1/n) \times 1 + \sum_{k=3}^n (1/n) \times 0 = 1/(n-1).$$

So, by Bayes' theorem,  $P(A_1 \text{ will be pardoned} \mid A \text{ is told that } A_3, A_4, \dots, A_n \text{ will be executed}) \\ = (1/n) \times [1/(n-1)] / [1/(n-1)] = 1/n.$

On the other hand,  $P(A_2 \text{ will be pardoned} \mid A \text{ is told that } A_3, A_4, \dots, A_n \text{ will be executed}) \\ = (1/n) \times 1 / [1/(n-1)] = (n-1)/n.$

**Assignment 4.** Denote by *Pos* and *Neg* the events that the test result is positive and negative, respectively. Also, denote by *D* and *ND* the events that the person actually has the disease and does not have it, respectively. We need to find  $P(D \mid Pos)$ . It is given that  $P(Pos \mid ND) = 1/100$ ,  $P(Neg \mid D) = 2/100$  and  $P(D) = 1/1000$ .

Now,  $P(Pos) = P(Pos \cap D) + P(Pos \cap ND)$

$$= P(D) - P(Neg \cap D) + P(Pos \cap ND)$$

$$= P(D) - P(D) \times P(Neg \mid D) + P(ND) \times P(Pos \mid ND)$$

$$= (1/1000) - (1/1000) \times (2/100) + (999/1000) \times (1/100) = 1097/(1000 \times 100) \approx 11/1000.$$

The above calculations show that  $P(Pos \cap D) = (1/1000) - (1/1000) \times (2/100) = 98/(1000 \times 100) \approx 1/1000$ .

So, we have  $P(D \mid Pos) = P(Pos \cap D) / P(Pos) = 98/1097 \approx 1/11$ .

Further,  $P(ND \mid Neg) = P(Neg \cap ND) / P(Neg)$

$$= [P(ND) - P(Pos \cap ND)] / [1 - P(Pos)]$$

$$= [(999/1000) - (999/1000) \times (1/100)] / [1 - 1097/(1000 \times 100)]$$

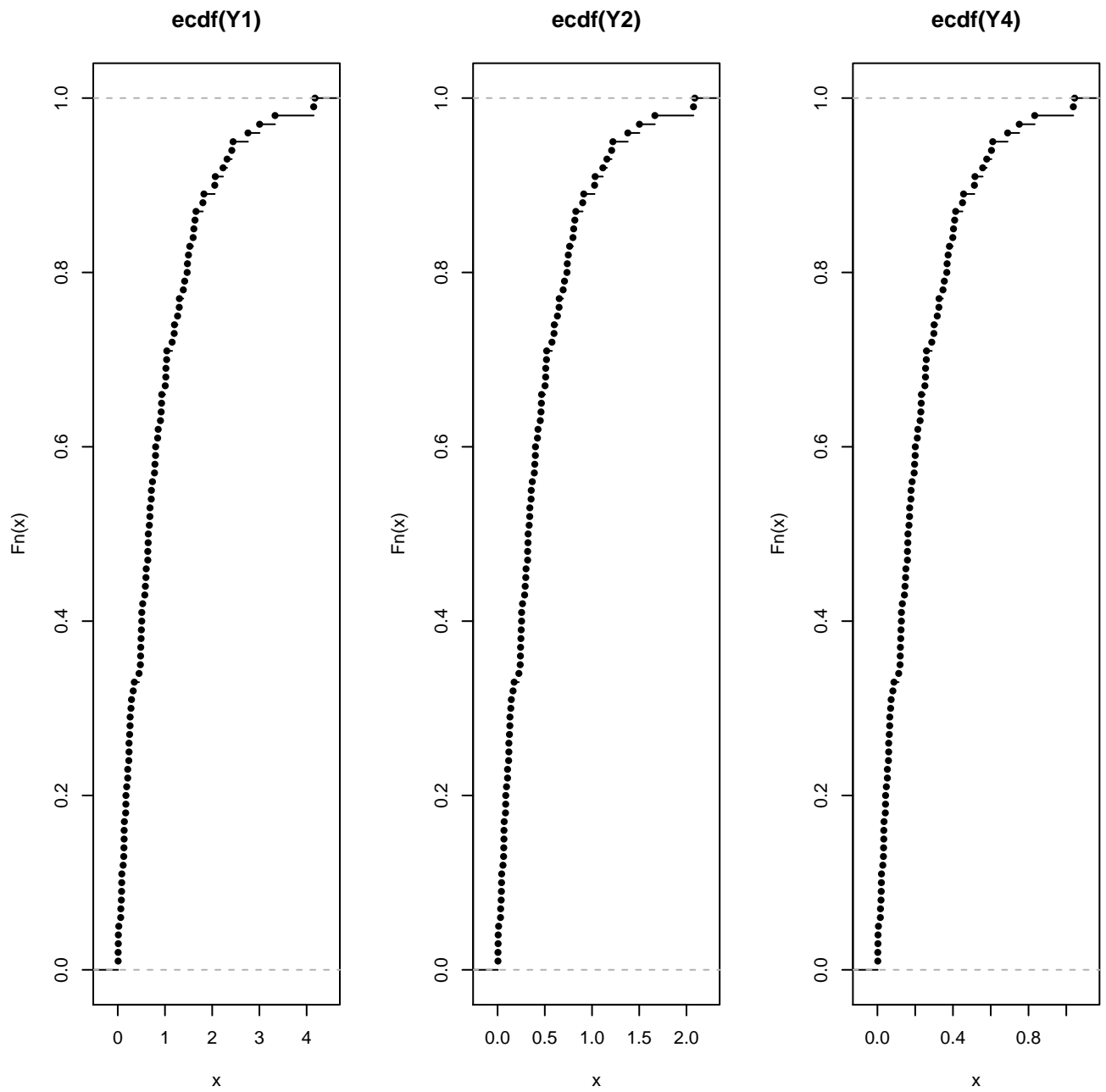
$$= 98901/98903 \approx 1.$$

Thus, although the chances of false positives (namely,  $P(Pos \mid ND)$ ) and true negatives (namely,  $P(Neg \mid D)$ ) from the test procedure are very small, it is not a good idea to base an affirmative diagnosis of the disease on this test alone.

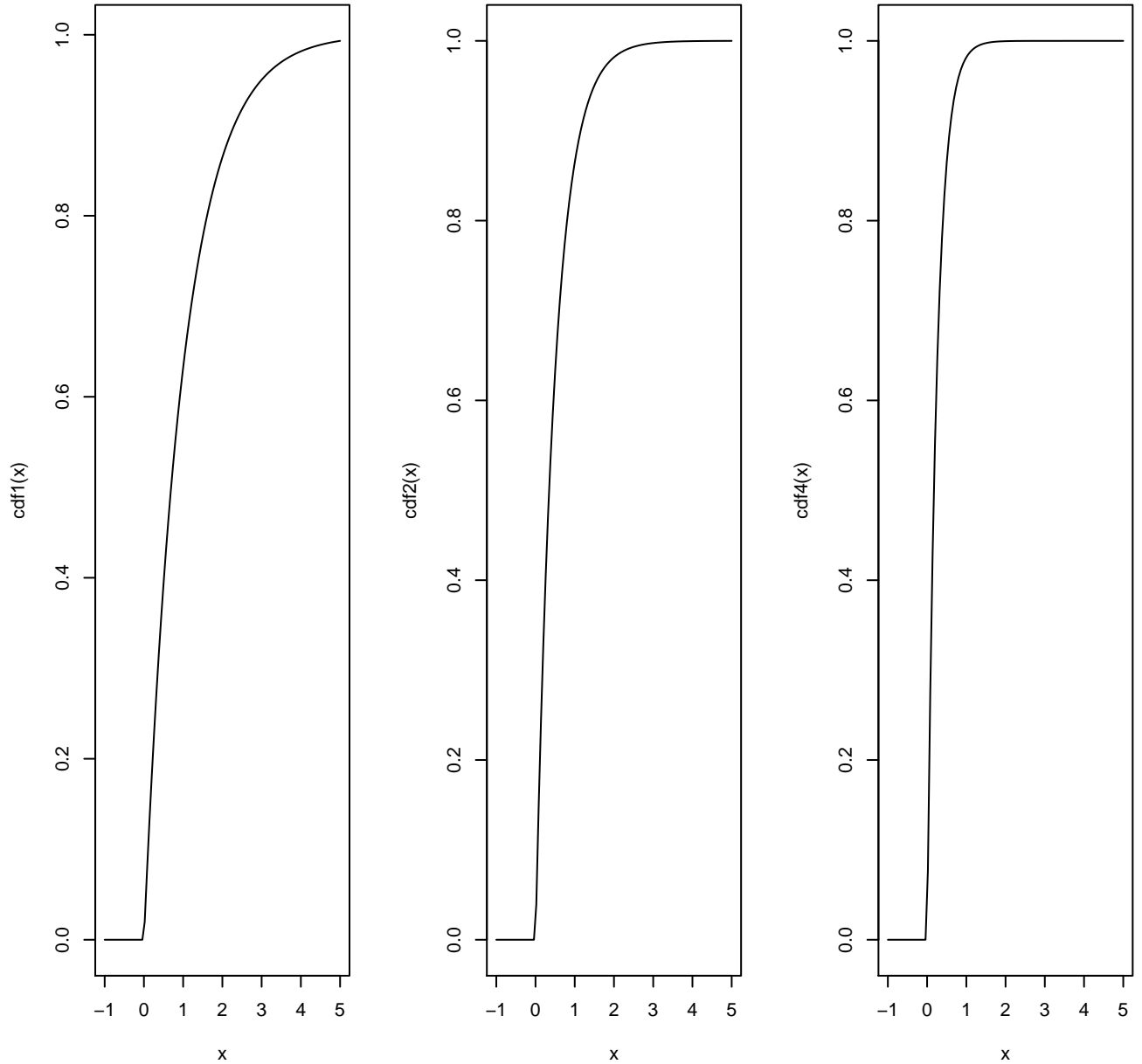
**Assignment 5.** (e) These plots show the empirical distribution functions

$$t \mapsto \frac{1}{100} \sum_{i=1}^{100} \mathbf{1}\{Y_i \leq t\}$$

for each of the vectors *Y1*, *Y2* and *Y4*. It is a nondecreasing right-continuous step function (piecewise constant) that equals 0 for *t* (strictly) smaller than the minimum of the sample, and 1 for *t* (weakly) larger than the maximum of the sample.



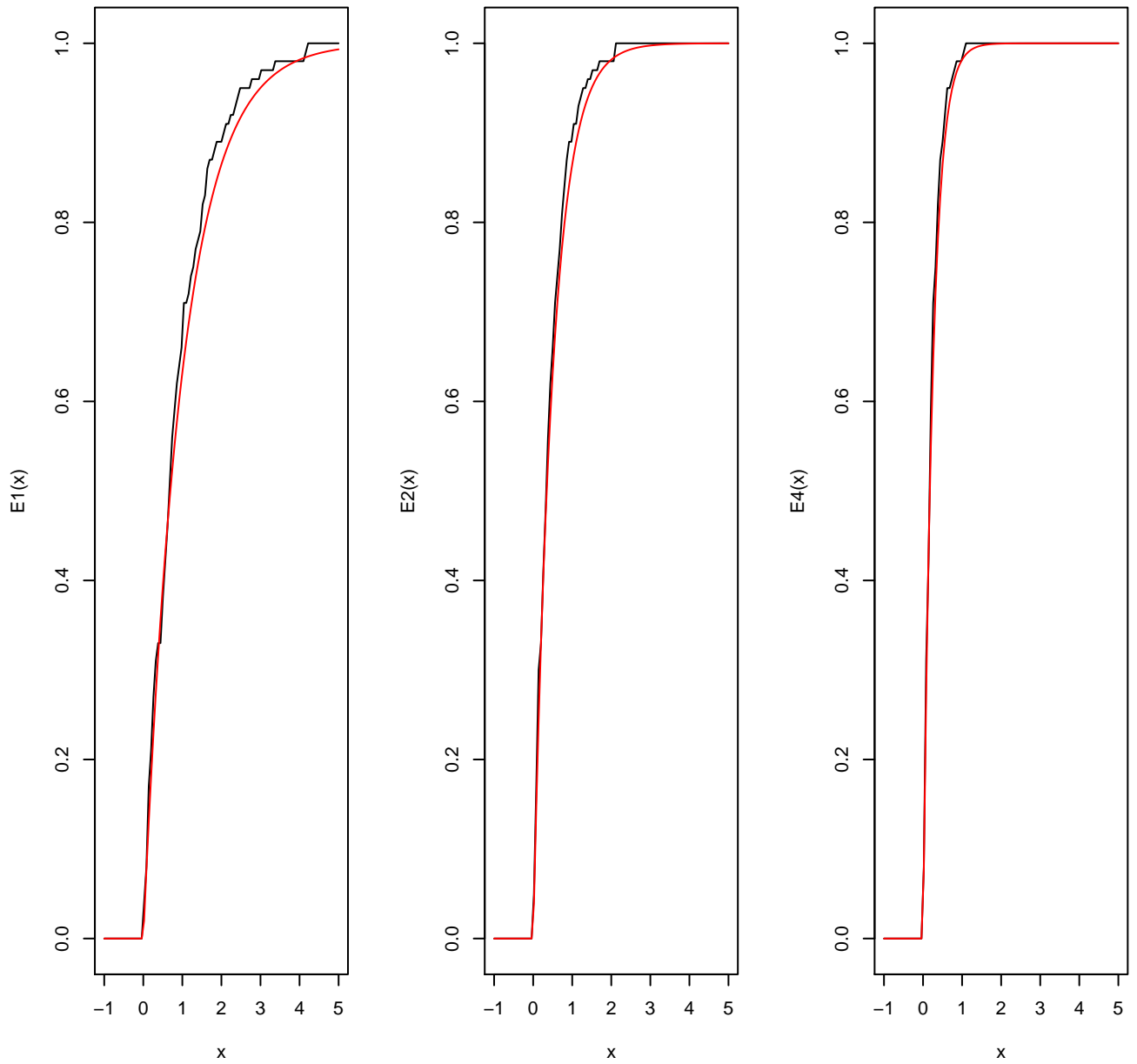
(g) These are plots of the functions  $F_\lambda$  for the parameter values 1, 2, and 4 for  $\lambda$ .



(i) These figures combine the two previous plots. We see that the empirical distribution function for  $Y\lambda$  is very close to the function  $F_\lambda$ . The figures imply that  $Y\lambda$  follows the distribution  $F_\lambda$ , i.e.  $\text{Exp}(\lambda)$ . And indeed, since  $X \sim \text{Unif}(0,1)$ ,

$$P(Y\lambda \leq t) = P(-\log(1 - X)/\lambda \leq t) = P(X \leq 1 - \exp(-\lambda t)) = F_\lambda(t).$$

The fact that the empirical distribution function approaches the distribution function itself is known as the *Glivenko–Cantelli* theorem.



(j)  $q_\lambda$  is the inverse of  $F_\lambda$  on  $(0, \infty)$ .

(k) Yes. This is so because  $Y = q(X) = F^{-1}(X) = F^{-1}(X)$  is such that

$$P(Y \leq t) = P(X \leq F(t)) = F(t).$$