

Errata for “Statistics for Data Science” Course

Updated – December 20, 2017

- Slide 27: If F_X is increasing **and continuous**, then $F_X^- = F_X^{-1}$.
- Slide 29: Partial Converse Theorem \rightarrow Let X be a random variable with **continuous** distribution F_X . Then, $F_X(X) \sim Unif(0, 1)$.
- Slide 39: The derivatives in the Jacobian matrix should be with respect to **y_i** and not x_i .
- Slide 40: X and Y should also be **independent**.
- Slide 40: The Jacobian matrix should be **transposed**; however this doesn't matter since we are interested in the determinant.
- Slide 42: Properties of Expectation $\rightarrow \mathbb{E}[h(X)]$ instead of $\mathbb{E}[h(x)]$.
- Slide 48: Definition of Conditional Expectation $\rightarrow \mathbb{E}[X | Y = y] = \sum_{x \in \mathcal{X}} x \mathbb{P}[X = x | Y = y]$, if X, Y are discrete.
- Slide 49: If X is independent of Y , then $\mathbb{E}[X | Y] = \mathbb{E}[X]$.
- Slide 52: All instances of \mathbb{R}^p should be **\mathbb{R}^d** .
- Slide 54: $M_{X+Y} = M_X M_Y$ **when X and Y are independent**.
- Slide 61: Lemma (Poisson and Multinomial) $\rightarrow \mathbf{X} = (X_1, \dots, X_k)^\top$.
- Slide 68: The inequalities $H(X) \geq 0$ and $H(g(X)) \leq H(X)$ hold when X is **discrete**.
- Slide 74:
 - “Natural” is from the mathematics point of view – usual parameter $\theta = \eta^{-1}(\phi)$ often different.
 - where $\eta : \mathbb{R}^k \rightarrow \mathbb{R}^k$ is a C^2 map such that $\phi = \eta(\theta)$.
- Slide 88: since T and **S** are 1-1 functions of each other.
- Slide 90:
 - without information loss **on** parameter of interest.
 - info can we **throw** away?
- Slide 91: Notice that for any \mathbf{y} , $\mathbf{w}_{T(\mathbf{y})}$ is in the same level set as **\mathbf{y}** .
- Slide 93: If the $\{T_j\}_{j=1}^k$ are non-trivial, the ratio $f(\mathbf{y})/f(\mathbf{z})$ will be constant **with respect to (ϕ_1, \dots, ϕ_k)** if and only if **as (ϕ_1, \dots, ϕ_k) varies, the quantity below remains constant**.
$$\sum_{j=1}^k \phi_j (\tau_j(y_1, \dots, y_n) - \tau_j(z_1, \dots, z_n)).$$
- Slide 102: its **form** may be tedious to work with

19. Slide 103:

(a) $Q_n = n(1 - M_n)$.

(b) $\mathbb{P}[Q_n \leq y] = \mathbb{P}[M_n \geq 1 - y/n] = 1 - (1 - y/n)^n$.

20. Slide 104: $T(Y_1, \dots, Y_n) \xrightarrow{d} Z$.

21. Slide 105:

(a) $\mathbb{P}[|Y_n - Y| > \epsilon] \xrightarrow{n \rightarrow \infty} 0$.

(b) **Example:** Let $U_1, \dots, U_n \stackrel{iid}{\sim} \mathcal{U}[0, 1]$ and $M_n = \max\{U_1, \dots, U_n\}$. Fix $\epsilon \in (0, 1)$.

$$\mathbb{P}[|M_n - 1| > \epsilon] = \mathbb{P}[M_n > 1 + \epsilon] + \mathbb{P}[M_n < 1 - \epsilon] = 0 + (1 - \epsilon)^n \xrightarrow{n \rightarrow \infty} 0.$$

Also, $\mathbb{P}[|M_n - 1| > \epsilon] = 0$ if $\epsilon \geq 1$. Hence, $M_n \xrightarrow{p} 1$ as $n \rightarrow \infty$.

22. Slide 107: $\mathbb{P}[Y_n \leq y] = \mathbb{P}[Y_n \leq y, |Y_n - Y| \leq \epsilon] + \mathbb{P}[Y_n \leq y, |Y_n - Y| > \epsilon]$.

23. Slide 112: Theorem (Central Limit Theorem) Let $\{Y_n\}$ be an i.i.d. sequence with mean μ and variance $\sigma^2 < \infty$. Then,

$$\sqrt{n} \left(\frac{1}{n} \sum_{i=1}^n Y_i - \mu \right) \xrightarrow{d} N(0, \sigma^2).$$

24. Slide 118:

(a) Theorem (Multivariate Law of Large Numbers) Let $\{\mathbf{Y}_n\}$ be iid random vectors with $\mathbb{E}[\mathbf{Y}_k] = \mu$ and $\mathbb{E}\|\mathbf{Y}_k\| < \infty$, for all $k \geq 1$. Then,

$$\frac{1}{n} \sum_{i=1}^n \mathbf{Y}_k \xrightarrow{p} \mu.$$

(b) Theorem (Multivariate CLT) Σ should be replaced by Ω .

25. Slide 133: $A^*\hat{\theta}$ should be replaced by $A^*(\theta)$.

26. Slide 135: $\mathbb{E}[\delta|T = t]$ in the first line of the proof should be replaced by $\mathbb{E}[\hat{\theta}|T = t]$.

27. Slide 150: If **both diagonal elements are positive**, then it will be positive definite.

28. Slide 151: The likelihood is zero if $\theta < Y_{(n)}$.

29. Slide 155: The latter is minimized at θ and so $\Psi(\mathbf{u})$ is maximized at θ .

30. Slide 156:

(a) Does $\{\Psi_n(\mathbf{u}) \xrightarrow{p} \Psi(\mathbf{u}) \forall \mathbf{u}$ with Ψ maximized uniquely at $\theta\} \dots$

(b) More general situations require stronger forms of convergence of $\Psi_n(\mathbf{u}) \rightarrow \Psi(\mathbf{u})$ plus additional regularity conditions.

31. Slide 159: In fact, the inverse function theorem tells us that the infinitely differentiable function $\Delta_\phi \gamma : \mathbb{R}^k \rightarrow \mathbb{R}^k$ must admit a continuously differentiable inverse map h locally at $\phi \in \mathbb{R}^k$.

32. Slide 162: Note that this can be interpreted as

$$\hat{\theta}_n \stackrel{d}{\approx} N \left(\theta, \frac{1}{n\mathcal{I}_1(\theta)} \right) \equiv N \left(\theta, \frac{1}{\mathcal{I}_n(\theta)} \right).$$

33. Slides 163 and 164: Why $\mathcal{I}_n(\theta)$? (... curvature)
34. Slide 173: $\tilde{\theta}_\alpha$ should be $\tilde{\mu}_\alpha$ and $R(\tilde{\mu}, \mu)$ should be $MSE(\hat{\mu}, \mu)$
35. Slide 180: all instances of \bar{X} should be \bar{Y}
36. Slide 211: $I^{-1}(\theta)$ should be $\mathcal{I}_1^{-1}(\theta_0)$
37. Slide 219: $\mathbb{P}_{H_0}[T^2(\mathbf{Y}) \geq T(\mathbf{y})]$ should be $\mathbb{P}_{H_0}[T^2(\mathbf{Y}) \geq T^2(\mathbf{y})]$
38. Slide 265: the first term in the last equation should be $\mathbb{E}[\int_{\mathbb{R}} \hat{f}_h^2(x) dx]$
39. Slide 266: the expectation in the first item should be $\mathbb{E}[\int_{\mathbb{R}} \hat{f}_h^2(x) dx]$, and its estimator is $\int_{\mathbb{R}} \hat{f}_h^2(x) dx$
40. Slide 266: the variables $\hat{f}_{h,-i}(Y_i)$ are **not** independent
41. Slide 267: the LSCV should be $\int_{\mathbb{R}} \hat{f}_h^2(x) dx - \frac{2}{n} \sum_{i=1}^n \hat{f}_{h,-i}(Y_i)$
42. Slide 273: Y is a random **output**
43. Slide 295: the vectors $\{v_j\}$ in the lemma are **not** unique
44. Slide 296: $v_i \Sigma v_i^T$ should be $v_i^T \Omega v_i$
45. Slide 297: $X^T X$ should be XX^T
46. Slide 346: the correct definitions of the adjusted R^2 are

$$R_a^2 = 1 - (1 - R^2) \frac{n-1}{n-p} \quad R_{0a}^2 = 1 - (1 - R_0^2) \frac{n}{n-p}.$$

47. Slide 446: Z_j should be $(W_j - 1\bar{W}_j)/(sd(W_j)\sqrt{n})$ (**divide** rather than multiply by \sqrt{n})
48. Slide 466: Y is a random **output**
49. Slide 470: $x_i \in \mathbb{R}^{p \times 1}$ so some transpositions are unnecessary:

$$\nabla_{\beta} \ell_n(\beta) = X_n^T Y - \sum_{i=1}^n x_i \gamma'(x_i^T \beta) = \sum_{i=1}^n x_i (Y_i - \mu_i) = X_n^T (Y - \mu)$$

50. Slide 472: the initialization should be $x_i^T \tilde{\beta} = (\gamma')^{-1}(Y_i)$